Performance and Reliability Analysis by Model Checking — From Verification to Synthesis —

Joost-Pieter Katoen



UNIVERSITY OF TWENTE.

7th oCPS Ph.D. School on Cyber-Physical Systems

Probabilistic Model Checking

The Relevance of Probabilities

Markov Models

Key Algorithms

Model Checking Fault Trees

Parameter Synthesis

Epilogue

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Overview

Probabilistic Model Checking



Probabilistic Model Checking

Gödel Prize 2000



Moshe Vardi



Pierre Wolper

"For work on model checking with finite automata."

Some other winners: Shor, Sénizergues, Agrawal et al., Spielman and Teng, ...

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Probabilistic Model Checking

ACM System Software Award 2001



Gerard J. Holzmann



SPIN book

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"SPIN is a popular open-source software tool, used by thousands of people worldwide, that can be used for the formal verification of distributed software systems."

Some other winners: TeX, Postscript, UNIX, TCP/IP, Java, Smalltalk, ...

ACM Turing Award 2007









E. Allen Emerson

Joseph Sifakis

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"For their role in developing model checking into a highly effective verification technology, widely adopted in the hardware and software industries."

Some other winners: Wirth, Dijkstra, Cook, Hoare, Rabin and Scott, ...

Striking Examples

- Needham-Schroeder protocol
- IEEE cache coherence protocol
- Hardware property languages like PSL
- C, .NET code verification
- NASA space mission software
- Storm surge barrier Maeslantkering



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Storm Surge Barrier Maeslantkering



Probabilistic Model Checking

Storm Surge Barrier Maeslantkering



10/131

Probabilistic Model Checking

Storm Surge Barrier Maeslantkering



What is This Lecture About?

"Probabilistic model checking is one of the main challenges for the future."



Edmund J. Clarke The Birth of Model Checking, 2008

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What is This Lecture About?

"A promising new direction in formal methods research these days is the development of probabilistic models, with associated tools for quantitative evaluation of system performance along with correctness."

Theory in Practice for System Design and Verification



Rajeev Alur Unix of Pennsylvania



Thomas A. Henzinger IST Austria



Moshe Y. Vardi Rice University

ACM SIGLOG News 2015

Overview

The Relevance of Probabilities

Almost Ten Reasons for Probabilities

- 1. Randomised Algorithms
- 2. Reducing Complexity
- 3. Avoiding the Impossible
- 4. Probabilistic Programs
- 5. Reliability
- 6. Performance
- 7. Robotics
- 8. Optimisation
- 9. Systems Biology

Randomised Algorithms: Simulating a Die [Knuth & Yao, 1976]



Heads = "go left"; tails = "go right". Does this model a six-sided die?

Avoiding the Impossible

FLP impossibility result

[Fischer et al., 1985]

In an asynchronous setting, where only one processor might crash, there is no distributed algorithm that solves the consensus problem—getting a distributed network of processors to agree on a common value.

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Ben-Or's possibility result

[Ben-Or, 1983]

If a process can make a decision based on its internal state, the message state, and some probabilistic state, consensus in an asynchronous setting is almost surely possible.

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The Relevance of Probabilities

Reliability Engineering

NUREG-0492

Fault Tree Handbook

U.S. Nuclear Regulatory Commission



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18/131

Reliability: Dynamic Fault Trees

[Dugan *et al.*, 1990]



The Relevance of Probabilities

A Fault Tree Example



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The Relevance of Probabilities

A Fault Tree Example



(D)FTs: one of —if not the— most prominent models for risk analysis Aims: quantify system reliability and availability, MTTF,

Performance: GSPNs

[Ajmone Marsan et al, 1984]

The early days:



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Performance: GSPNs

[Ajmone Marsan et al, 1984]

The early days:



More modern times: Petri nets with

- Timed transitions
- Immediate transitions
- Natural weights



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Performance: GSPNs

[Ajmone Marsan et al, 1984]

The early days:



More modern times: Petri nets with

- Timed transitions
- Immediate transitions
- Natural weights



Aims: quantify arrivals, waiting times, QoS, soft deadlines, GSPNs: very —if not the most— popular in performance modeling

Encyclopedia of Optimisation 2008

Stochastic Scheduling

JOSÉ NIÑO-MORA Department of Statistics, Universidad Carlos III de Madrid, Getafe, Spain

MSC2000: 90B36

Article Outline

Introduction Models Scheduling a Batch of Stochastic Jobs Multi-Armed Bandits Scheduling Queueing Systems References

Introduction

The field of stochastic scheduling is motivated by problems of priority assignment arising in a variety of systems where jobs with random features (e.g. arrival or

optimal performance. The theory of stoc a goal in the idealized els. Real-world rando rival or processing tir ing their probability (vary across several d scheduling policies cc terarrival and proces arrangement of servic jective to be optimize are required to be nc cannot make use of fu known total duration vet finished.

Regarding solutic seems fair to say that : is yet available to des optimal policies acrc tic scheduling model can be cast in the fi

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22/131

Stochastic Scheduling

[Bruno et al., 1981]

- Job processing times are subject to random variability
 - machine breakdowns and repairs, job parameters, ...
 - N independent jobs with mean duration $\frac{1}{\mu_i}$
 - M identical machines
 - job processing with (or without) pre-emption
- Objective = minimal expected makespan, i.e., finishing time of last job
- SEPT policy yields minimal expected makespan
 "it is hard to calculate these expected values"

Which policy maximises the probability to finish all jobs on time?

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The Relevance of Probabilities

Stochastic Model



Nature Behaviour: Systems Biology

Enzyme-catalysed substrate conversion

 W http://en.wikipedia.org/wiki/Enzyme		
reaction, the reaction is effectively irreversible. Under these con catalyze the reaction in the thermodynamically allowed direction Kinetics	ditions the enzyme will, in fact, only n.	stabilizes the transition state, form this species and thus rea form products.
Catalytic step Enzyr E + S + ES E + P Substrate binding Substrate binding Methodshism for a single substrate enzyme catalyzed Substrate Methodshism for a single substrate enzyme (E) binds a substrate (S) and produces Menh a product (P). Substrate complex. This is sometimes called the Micha product.	Enzyme kinetics is the investigation of how enzymes bind substrates and tu rate data used in kinetic analyses are obtained from enzyme assays. In 1902 Victor Henri ^[45] proposed a quantitative theory of enzyme kinetics, were not useful because the significance of the hydrogen ion concentration After Peter Launtz Sorensen had defined the logarithmic pH-scale and intro buffering in 1909 ^[46] the German chemist Leionor Michaelis and his Canadi Menten repeated Henri's experiments and confirmed his equation which is Henri-Michaelis-Menten kinetics (sometimes also Michaelis-Menten kinetic developed by G. E. Briggs and J. B. S. Haldane, who derived kinetic equa used today. ^[46]	

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Stochastic Chemical Kinetics

Types of reaction described by stochiometric equations:

$$E + S \stackrel{k_1}{\rightleftharpoons} C \stackrel{k_3}{\to} E + P$$

► N different types of molecules that randomly collide where state X(t) = (x₁,...,x_N) with x_i = # molecules of sort i

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Stochastic Chemical Kinetics

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- ► N different types of molecules that randomly collide where state X(t) = (x₁,...,x_N) with x_i = # molecules of sort i
- Reaction probability within infinitesimal interval $[t, t+\Delta)$: $\alpha_m(\vec{x}) \cdot \Delta = \Pr\{\text{reaction } m \text{ in } [t, t+\Delta) \mid X(t) = \vec{x}\}$

where $\alpha_m(\vec{x}) = k_m \cdot \#$ possible combinations of reactant molecules in \vec{x}

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Stochastic Chemical Kinetics

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- ► N different types of molecules that randomly collide where state X(t) = (x₁,...,x_N) with x_i = # molecules of sort i
- ► Reaction probability within infinitesimal interval [t, t+Δ): α_m(x) · Δ = Pr{reaction m in [t, t+Δ) | X(t) = x} where α_m(x) = k_m · # possible combinations of reactant molecules in x
- Process has the Markov property and is time-homogeneous

Substrate Conversion in the Small



Transitions:
$$E + S \stackrel{1}{\rightleftharpoons} C \stackrel{0.001}{\rightarrow} E + P$$

e.g., $(x_E, x_S, x_C, x_P) \stackrel{0.001 \cdot x_C}{\rightarrow} (x_E + 1, x_S, x_C - 1, x_P + 1)$ for $x_C > 0$

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7/131

Overview

Markov Models

Common Feature

All these applications consider Markov models¹

¹Non-exponential distributions are approximated by phase-type distributions. Ξ

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29/131

Discrete-Time Markov Models



A Markov chain for Knuth-Yao's algorithm



A Markov decision process for a probabilistic program

30/131
Markov Models

Continuous-Time Markov Models



A Markov chain for substrate conversion



A Markov decision process for the GSPN

Markov Models

Continuous-Time Markov Models



Markov decision process for stochastic scheduling



Markov decision process for a DFT

	Discrete	Continuous
Deterministic	discrete-time Markov chain (DTMC)	continuous-time MC
Nondeterministic	Markov decision process (MDP)	CTMDP
Compositional	Segala's probabilistic automata (PA)	Markov automata (MA)

Other models: e.g., probabilistic timed automata, pVASS, pPDA, SGs, etc.

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Overview

Key Algorithms

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	Discrete	Continuous
Logic	probabilistic CTL	probabilistic timed CTL
Monitors	deterministic automata (safety and LTL)	deterministic timed automata (MITL fragments)

Others: e.g., conditional probs, multi-objective, rewards, quantiles, etc.

Core problem: computing (timed) reachability probabilities

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Problem

Consider a finite MC with $s \in S$ and $G \subseteq S$.

Aim: determine $\Pr(s \models \diamondsuit G) = \Pr_s \{ \pi \in Paths(s) \mid \pi \models \diamondsuit G \}$

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• Let variable $x_s = \Pr(s \models \diamondsuit G)$ for any state s

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 - if **G** is not reachable from s, then $x_s = 0$
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- For any state $s \in Pre^*(G) \setminus G$:

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- For any state $s \in Pre^*(G) \setminus G$:

$$x_{s} = \underbrace{\sum_{t \in S \setminus G} \mathbf{P}(s, t) \cdot x_{t}}_{\text{reach } G \text{ via } t \in S \setminus G} + \underbrace{\sum_{u \in G} \mathbf{P}(s, u)}_{\text{reach } G \text{ in one step}}$$

Reachability Probabilities: Knuth-Yao's Die

Consider the event 4





- Consider the event
- We obtain:

$$x_1 = x_2 = x_3 = x_5 = x_6 = 0$$
 and $x_4 = 1$



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- Consider the event <>4
- We obtain:

$$x_{1} = x_{2} = x_{3} = x_{5} = x_{6} = 0 \text{ and } x_{4} = 1$$
$$x_{s_{1}} = x_{s_{3}} = x_{s_{4}} = 0$$
$$x_{s_{0}} = \frac{1}{2}x_{s_{1}} + \frac{1}{2}x_{s_{2}}$$



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$$x_{s_{5}} = \frac{1}{2}x_{5} + \frac{1}{2}x_{4}$$

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Reachability Probabilities: Knuth-Yao's Die



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$$x_{s_{5}} = \frac{1}{2}x_{5} + \frac{1}{2}x_{4}$$

$$x_{s_{6}} = \frac{1}{2}x_{s_{2}} + \frac{1}{2}x_{6}$$

Gaussian elimination yields:

$$x_{s_5} = \frac{1}{2}, x_{s_2} = \frac{1}{3}, x_{s_6} = \frac{1}{6}, \text{ and } x_{s_0} = \frac{1}{6}$$

• Repeated reachability $Pr(s \models \Box \diamondsuit G)$:

Determine probability to reach a terminal SCCs containing a G-state

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Probabilistic CTL model checking

Recursive descent on parse tree using reach-probabilities at nodes

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Probabilistic CTL model checking

Recursive descent on parse tree using reach-probabilities at nodes

- LTL formulas $Pr(s \models \varphi)$:
 - 1. Transform φ into a deterministic (Rabin) automaton
 - 2. Take the product of the Markov chain and the automaton
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For MDPs, solving linear inequality systems are key.

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Randomised Algorithms: Simulating a Die [Knuth & Yao, 1976]



Probability of after initial tails, yield 1 or 3 but with at most five tails in total?

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39/131

Property of Knuth-Yao's Algorithm



After initial tails, yield 1 or 3 but with at most five times tails in total

Product Markov Chain



Reachability probability of terminal SCC with (\cdot, q_{acc}) is $\frac{1}{8} + \frac{1}{8} + \frac{1}{32} + \frac{1}{32} = \frac{5}{16}$.

Probabilistic CTL

Key Algorithms

[Hansson & Jonsson, 1989]

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Performance+Reliability by Model Checking 42/1

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[Hansson & Jonsson, 1989]

• PCTL interpretation is Boolean, i.e., a formula is satisfied or not.

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- PCTL interpretation is Boolean, i.e., a formula is satisfied or not.
- For path-formula φ and threshold $\succ p$ with $\succ \in \{>, \ge\}$ and $p \in \mathbb{Q}$:

PCTL-formula $[\varphi]_{\succ p}$ denotes

all paths satisfying φ occur with probability $\succ p$

• $[\cdot]_{>p}$ is probabilistic counterpart of CTL path-quantifiers \exists and \forall .

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- Examples: $[\diamondsuit a]_{>1/2}$, $[\diamondsuit [\Box a]_{=1}]_{>1/2}$ and $[\Box (\neg a \land [\diamondsuit a]_{>0})]_{>0}$.

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PCTL model checking is in P.

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CTL versus Probabilistic CTL

- ▶ Qualitative PCTL (only >0- bounds) and CTL are incomparable.
 - 1. There is no CTL formula that is equivalent to $[\diamondsuit a]_{=1}$.
 - 2. There is no PCTL formula that is equivalent to $\forall \diamondsuit a$.
- These results all rely on countably infinite MCs
 - 1. For finite MCs, $[\diamondsuit a]_{=1} \equiv \forall \diamondsuit a$ under fairness.
 - 2. For finite MCs, *⇔*-modalities are PCTL-definable, but not in CTL.

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Random Timing

Key Algorithms





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Continuous-Time Markov Chains

A CTMC is a DTMC with an *exit rate* function $r: S \to \mathbb{R}_{>0}$ where r(s) is the rate of an exponential distribution.



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Zeno theorem

In every CTMC, almost surely no Zeno runs occur.

In contrast to timed automata verification, Zeno runs thus pose no problem.

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Timed Reachability Probabilities

Problem

Consider a finite CTMC with $s \in S$, $t \in \mathbb{R}_{\geq 0}$ and $G \subseteq S$.

Aim: determine $\Pr(s \models \diamondsuit^{\leq t} G)$.

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Characterisation of timed reachability probabilities

• Let function $x_s(t) = \Pr(s \models \diamondsuit^{\leq t} G)$ for any state s

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- For any state $s \in Pre^*(G) \setminus G$:

$$x_{s}(t) = \int_{0}^{t} \sum_{s' \in S} \underbrace{\mathsf{R}(s, s') \cdot e^{-r(s) \cdot x}}_{\text{probability to move to}} \cdot \underbrace{x_{s'}(t-x)}_{\text{prob. to fulfill}} dx$$

$$state s' \text{ at time } x \qquad \diamondsuit^{\leqslant t-x} G \text{ from } s'$$



Integral equations for $\diamondsuit^{\leq 10} 2$:

•
$$x_3(d) = 0$$
 and $x_2(d) = 1$ for all d
• $x_0(d) = \int_0^d \frac{25}{4 \cdot e^{-25 \cdot x}} \cdot x_1(d-x) + \frac{25}{4 \cdot e^{-25 \cdot x}} \cdot x_2(d-x) dx$
• $x_1(d) = \int_0^d \frac{4}{2 \cdot e^{-4 \cdot x}} \cdot x_0(d-x) + \frac{4}{2 \cdot e^{-4 \cdot x}} \cdot x_3(d-x) dx$

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Reachability probabilities

Solve a system of linear equations for which many efficient techniques exist.

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Reachability probabilities

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Timed reachability probabilities

Solve a system of Volterra integral equations. Non-trivial, inefficient, and has several pitfalls such as numerical stability.

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Reachability probabilities

Solve a system of linear equations for which many efficient techniques exist.

Timed reachability probabilities

Solve a system of Volterra integral equations. Non-trivial, inefficient, and has several pitfalls such as numerical stability.

Solution

Reduce $\Pr(s \models \diamondsuit^{\leq t} G)$ to computing transient probabilities.

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Timed Reachability Probabilities = Transient Probabilities

Aim

Compute $\Pr(s \models \diamondsuit^{\leq t} G)$ in CTMC C. Observe that once a path π reaches G within t time, then the remaining behaviour along π is not important. \Rightarrow make all states in G absorbing.

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Timed Reachability Probabilities = Transient Probabilities

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$$\underbrace{\Pr(s \models \diamondsuit^{\leq t} G)}_{\text{timed reachability in } \mathcal{C}} = \underbrace{\Pr(s \models \diamondsuit^{=t} G)}_{\text{timed reachability in } \mathcal{C}[G]} = \underbrace{\vec{p}(t) \text{ with } \vec{p}(0) = \mathbf{1}_s}_{\text{transient prob. in } \mathcal{C}[G]}.$$

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Timed Reachability Probabilities = Transient Probabilities

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Transient probabilities can be efficiently computed as solutions of linear differential equations.

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Computing Transient Probabilities

By solving a linear differential equation system

The transient probability vector $\underline{p}(t) = (p_{s_1}(t), \dots, p_{s_k}(t))$ satisfies:

$$\underline{p}'(t) = \underline{p}(t) \cdot (\mathbf{R} - \mathbf{r})$$
 given $\underline{p}(0)$

where \mathbf{r} is the diagonal matrix of vector \underline{r} .

 2 19 dubious ways to compute a matrix exponential [Moler & Van Loan, 1978/2003].

Computing Transient Probabilities

By solving a linear differential equation system

The transient probability vector $\underline{p}(t) = (p_{s_1}(t), \dots, p_{s_k}(t))$ satisfies:

$$\underline{p}'(t) = \underline{p}(t) \cdot (\mathbf{R} - \mathbf{r})$$
 given $\underline{p}(0)$

where \mathbf{r} is the diagonal matrix of vector \underline{r} .

Solution using standard knowledge yields: $\underline{p}(t) = \underline{p}(0) \cdot e^{(\mathbf{R}-\mathbf{r}) \cdot t}$.

 2 19 dubious ways to compute a matrix exponential [Moler & Van Loan, 1978/2003].

51/131

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Computing the matrix exponential is a challenging numerical problem².

²19 dubious ways to compute a matrix exponential [Moler & Van Loan, 1978/2003].

CTMC C is uniform if r(s) = r for all $s \in S$ for some $r \in \mathbb{R}_{>0}$.

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Uniformisation	[Jensen, 1953]	[Gross and Miller, 1984]
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$\overline{\mathbf{P}}(s,s') = \frac{r(s)}{r} \cdot \mathbf{P}(s,s') \text{ if } s' \neq s \text{ and}$	$\overline{P}(s,s) = \frac{r(s)}{r} \cdot P(s,s) + 1 - \frac{r(s)}{r}.$

 $\overline{\mathbf{P}}$ is a stochastic matrix and $\overline{r}(\mathcal{C})$ is uniform.

Uniformisation by Example



Uniformisation amounts to normalise the residence time in every CTMC state.

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Benefits of Uniformisation

Transient probabilities of a CTMC and its uniformized CTMC coincide.

Thus:
$$\underline{p}(t) = \underline{p}(0) \cdot e^{(\mathbf{R} - \mathbf{r}) \cdot t} =$$

transient probablity in $\ensuremath{\mathcal{C}}$

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Still a matrix exponential remains. Did we gain anything? Yes. Since $\overline{\mathbf{P}}$ is stochastic, Taylor-Maclaurin yields $\sum_{i} \dots \overline{\mathbf{P}}^{i}$.

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Expected time objectives

Can be characterised as solution of set of linear equations

Expected time objectives

Can be characterised as solution of set of linear equations

- Long-run average objectives
 - 1. Determine the limiting distribution in any terminal SCC
 - 2. Take weighted sum with reachability probabilities terminal SCCs

³This yields a piecewise deterministic Markov process. (=) (=

Expected time objectives

Can be characterised as solution of set of linear equations

- Long-run average objectives
 - 1. Determine the limiting distribution in any terminal SCC
 - 2. Take weighted sum with reachability probabilities terminal SCCs
- Probabilistic timed CTL model checking

recursive descent over parse tree

Expected time objectives

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recursive descent over parse tree

Deterministic timed automata objectives

- 1. Take product of the MC and the Zone automaton of the DTA³
- 2. Determine the probability to reach an accepting zone

Timed Reachability in CTMDPs is Hard





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Timed Reachability in CTMDPs is Hard



- Timed policies are optimal; any time-abstract policy is inferior.
- If long time remains: choose β ; if short time remains: choose α .
- Optimal for deadline 1: choose α if $1-t_0 \leq \ln 3 \ln 2$, otherwise β

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Probabilistic Model Checkers

► PRISM ⁴	[Kwiatkowska, Parker <i>et al</i> .]
► MRMC	[Katoen <i>et al.</i>]
▶ iscasMC	[Zhang <i>et al.</i>]
▶ iBioSim	[Myers <i>et al.</i>]
► GreatSPN	[Franceschinis <i>et al.</i>]
► SMART	[Ciardo <i>et al.</i>]
▶ MarCie	[Heiner <i>et al.</i>]
► PAT	[Song Dong <i>et al.</i>]
► SToRM	[Dehnert, Katoen <i>et al.</i>]

Statistical model checkers: Ymer, Vesta, UppAal-SMC, PlasmaLab,

⁴Recipient HVC Award 2016.

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The Probabilistic Model Checker SToRM



Supports Markov chains, CTMCs, MDPs, and CTMDPs About 100,000 lines of C++ code

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Comparison to PRISM



More information at: stormchecker.org

Experimental Comparison



Comparing the best engines for all

Overview

- Model Checking Fault Trees

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Model Checking Fault Trees

Fault Tree Analysis



Dynamic Fault Trees

[Dugan *et al.*, 1995]



Markov decision process for a DFT

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Experiences with FT model checking







- Logics like PCTL allow for expressing more properties
 - but hide logics as far as possible: use specification patterns or so
- Enable the analysis of a larger class of DFTs
- Model checking mostly substantially faster than FT analysis
- Abstraction aggravates this-for traditional FTA-even further:
 - 1. compositional minimisation
 - 2. tailored abstractions for FTs
 - 3. symmetry reduction and modularisation on FTs
 - 4. aggressive abstraction on FTs

yield several orders of magnitude improvements.

Model Checking Fault Trees

Probabilistic Bisimulation

Intuition: transition probabilities for each equivalence class coincide.

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Probabilistic bisimulation

[Larsen & Skou, 1989]

Consider a DTMC with state space S and equivalence $R \subseteq S \times S$. R is a probabilistic bisimulation on S if for any $(s, t) \in R$:

$$L(s) = L(t)$$
 and $P(s, C) = P(t, C)$ for each $C \in S/R$

where $P(s, C) = \sum_{s' \in C} P(s, s')$.

Let ~ denote the largest possible probabilistic bisimulation.

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Probabilistic bisimulation

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Let \sim denote the largest possible probabilistic bisimulation.

Variants: weak, divergence-sensitive, distribution-based, for CTMC, MDPs, etc.

Model Checking Fault Trees

Craps

- Come-out roll:
 - ▶ 7 or 11: win
 - 2, 3, or 12: lose
 - else: roll again
- Next roll(s):
 - 7: lose
 - point: win
 - else: roll again



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Craps's Bisimulation Quotient



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Quotienting: using partition-refinement in \mathcal{O}(|\mathbf{P}| \cdot \log |S|)
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Preservation: all probabilistic CTL*-formulas

Congruence: with respect to parallel composition

 $(\mathcal{M}_1 \sim \mathcal{N}_1 \text{ and } \mathcal{M}_2 \sim \mathcal{N}_2)$ implies $\mathcal{M}_1 \| \mathcal{M}_2 \sim \mathcal{N}_1 \| \mathcal{N}_2$

Stuttering: weak variants preserve PCTL* without next-modalities

Savings: potentially exponentially in time and space

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Exploiting Compositionality

Assume system is given by:

$$\mathcal{M}_1 \| \dots \| \mathcal{M}_i \| \dots \| \mathcal{M}_k$$

with \mathcal{M}_i a Markov automaton and CSP-like composition ||

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Exploiting Compositionality

[Hermanns and K., 2000]

• Assume system is given by:

$$\mathcal{M}_1 \| \dots \| \mathcal{M}_i \| \dots \| \mathcal{M}_k$$

with \mathcal{M}_j a Markov automaton and CSP-like composition \parallel

• Recall congruence property: $(\mathcal{M}_1 \sim \mathcal{N}_1 \text{ and } \mathcal{M}_2 \sim \mathcal{N}_2) \text{ implies } \mathcal{M}_1 \| \mathcal{M}_2 \sim \mathcal{N}_1 \| \mathcal{N}_2$

Exploiting Compositionality

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- Component-wise minimisation
 - 1. Pick process \mathcal{M}_i and consider its quotient \mathcal{M}_i/\sim under \sim
 - 2. Yielding $\mathcal{M}_1 \parallel \ldots \parallel \mathcal{M}_i / \sim \parallel \ldots \parallel \mathcal{M}_k$; repeat 1. and 2.
 - 3. Once all done, minimise pairs $\mathcal{M}_i/\sim ||\mathcal{M}_{i+1}/\sim \text{etc.}$

[Crouzen et al., 2010]





[Crouzen et al., 2010]



CPS 133 465 .00135 67

Comparing Galileo DIFTree (top) to DFTCalc (bottom)

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[Crouzen et al., 2010]



CPS	133	465	.00135	67
CAS	36	119	.65790	94

Comparing Galileo DIFTree (top) to DFTCalc (bottom)

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[Crouzen et al., 2010]

case study	peak # sti	* ^{tra}	nsitions	urreliability	time
CPS	4113	3 24	4608	.00135	490
CAS	8		10	.65790	1
CAS-PH	х		х	х	х
NDPS	х		х	х	х
CPS	133	465		0013	5 67
CAS	36	119		.6579	0 94
CAS-PH	40052	265442		.11	2 231
NDPS	61	169	[.00	586, .00598	266

Comparing Galileo DIFTree (top) to DFTCalc (bottom)

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[Crouzen et al., 2010]

	case study		peak # sta	^{رو5} * ^{۲۵۲}	Sitions	unreliabilit	y time	رہ	
-	CP	PS 4113		24608		.0013	54	90	
	CA	S	8		10	.6579	0	1	
	CAS-PH		х		х	2	x	х	
	NDPS		х		х	2	x	х	
	FTTP-4		32757	7 426	6826	.01922	2 131	11	
	FTTP-5		МО		мо	м	о і	мо	
C	PS	13	33	46	5		.00135	67	7
CAS 3		6	11	9		.65790	94	ł	
CAS	CAS-PH 40052)52	26544	2		.112	231	L
NE	IDPS 61		16	169 [.005		.00598]	266	5	
FT	TP-4 1325		1364	13642		.01922	65	5	
FTTP-6 11806		6565	2214737	8		.00045	1989)	

Comparing Galileo DIFTree (top) to DFTCalc (bottom)

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Tailored DFT Abstraction

[Junges et al., 2015]

Key idea

Simplify DFTs by graph rewriting prior to (compositional) state space generation.



Tailored DFT Abstraction



total verification and minimisation timestate space size of resulting CTMDP49 out of 179 case studies could be treated now that could not be treated before

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Boosting DFT State Space Generation

[Volk et al., 2016]

Apply POR, symmetry reduction, bisimulation, and state bit vectors.



 \Rightarrow This boosts FT analysis by several orders of magnitude.

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- Partition the state space into groups of concrete states
 - allow any partitioning, not just grouping of bisimilar states

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- Partition the state space into groups of concrete states
 - allow any partitioning, not just grouping of bisimilar states
- This typically yields over-approximations
 - abstraction yields safe bounds on true measures
- Correctness relies on simulation relations
 - preserve safety fragments of PCTL
- Various abstract probabilistic models exist
 - two-player SGs, interval MCs, abstract PA, modal models, etc.

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[Kwiatkowska et al., 2010]



Millions of states can be reduced to hundreds of states in a few AR iterations

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Abstraction-Refinement for DFTs

[Volk et al., 2017]

Partial fault tree analysis, making best/worst-case assumptions.



Abstraction-refinement terminates at 10% precision: $u-\ell < 1/10 \cdot \frac{u-\ell}{2}$ \Rightarrow Scalable and one order of magnitude faster.

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Model Checking Fault Trees

Safety Analysis of Vehicle Guidance







Major safety goal: avoid wrong vehicle guidance.

Fail Operational





Automotive Safety Integrity Level (ASIL)

Model Checking Fault Trees





Vehicle guidance ASIL-D: 10⁻⁸ residual HW failures/hour

Approach

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Vehicle Guidance





Fail-operational design patterns for autonomous driving.

Modern Car Architectures





DFT Statistics



	Scenario					DFT			CTMC	CTMC	
	SC	Arch.	Adap.	Sens.	Act.	#BE	#Dyn.	#Elem.	#States	#Trans.	Degrad.
Ι	SC1	В	-	2/4	4/4	76	25	233	5,377	42,753	-
П	SC2	В		2/4	4/4	70	23	211	5,953	50,049	19.35%
III	SC2	C	ADAS+	2/4	4/4	57	19	168	1,153	7,681	16.65%
IV	SC3	C	-	2/4	4/4	57	21	170	385	1,985	12.47%
V	SC2	A		2/4	4/4	58	19	185	193	897	0.00%
VI	SC2	в	removed I-ECU	2/4	4/4	65	21	199	1,201	8,241	19.98%
VII	SC2	в	5 ADAS, 2 BUS	2/8	7/7	96	30	266	194,433	2,171,905	19.35%
VIII	SC2	B	8 ADAS, 2 BUS	6/8	7/7	114	36	305	3,945,985	66,225,665	10.90%

One of the largest real-life DFTs in the literature

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Model Checking Fault Trees

Analysis Approach



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	Measure	Model Checking Queries
System	integrity	$1 - P(\Diamond^{\leq t} \text{ failed})$
	FIT	$\frac{1}{\text{lifetime}} \cdot \left(1 - P(\Diamond^{\leq \text{lifetime}} \text{ failed})\right)$
	MTTF	ET(◊ failed)
Degradation	FFA	$1 - P(\Diamond^{\leq t} \text{ (failed } \lor \text{degraded}))$
	FWD	$P((\neg \text{degraded}) \ U^{\leq t} \ (\neg \text{degraded} \land \text{failed}))$
	MTDF	$\Sigma_{s \in \text{degraded}} \left(P(\neg \text{degraded } U s) \cdot ET^s(\Diamond \text{ failed}) \right)$
	MDR	$\operatorname{argmin}_{s \in \operatorname{degraded}} \left(1 - P^s(\Diamond^{\leq t} \text{ failed}) \right)$
	SILFO	$1 - \left(FWD + \Sigma_{s \in \text{degraded}} \left(P(\neg \text{degraded } \bigcup^{\leq t} s) \cdot P^s(\Diamond^{\leq \text{drivecycle failed}}) \right) \right)$

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Model Checking Results



Model Checking Fault Trees



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Joost-Pieter Katoen

Overview

- Parameter Synthesis

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The Need for Parameter Synthesis

Fact:

Probabilistic model checking is applicable to various areas, e.g.:

- fault trees
- randomised algorithms
- systems biology

Limitation:

Probabilities need to be known a priori. Is this a valid assumption? How sensitive are results when transition probabilities fluctuate?

Goal:

Treat parametric models, synthesise "safe" parameter values

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Parameter Synthesis

Biased Knuth-Yao's Die



Joost-Pieter Katoen

Performance+Reliability by Model Checking

Parameter Synthesis

Biased Knuth-Yao's Die



Joost-Pieter Katoen

Performance+Reliability by Model Checking

Parameter Synthesis

Parametric Knuth-Yao's Die



For which $1/10 \le p \le 9/10$ and $2/5 \le q \le 3/5$ does $\Pr(\diamondsuit 2) \ge 3/20$ hold?
Conditional Probabilities



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Inputs:

- 1. a (finite) parametric Markov model
- 2. a property (e.g., reachability, expected reward, conditional reachability)
- 3. a threshold

Output:

For which parameter values does the pMC satisfy the property with the given threshold?

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Inputs:

- 1. a (finite) parametric Markov model
- 2. a property (e.g., reachability, expected reward, conditional reachability)
- 3. a threshold

Output:

For which parameter values does the pMC satisfy the property with the given threshold?

Problem instances:

- What is the maximal tolerable message loss ensuring delivery $\ge 98\%$?
- ... the tolerable failure rate in a DFT ensuring MTTF ≥ 3 hours?

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Recall Dynamic Fault Trees



Markov chain process for a DFT

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Parametric Fault Trees



Sample parametric DFT and its MTTF

$$\mathsf{MTTF} = \frac{200x^2 + 20x + 201}{x \cdot (20x + 201)} \text{ for } (\alpha, \beta, \gamma, d) = (10, x, 0.1, 0.5)$$

Joost-Pieter Katoen

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Parameter Synthesis

Aim:

partition the parameter space into safe and unsafe regions

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Aim:

partition the parameter space into safe and unsafe regions

- ▶ Region = half-space defined by linear inequalities over the parameters
- A region R for threshold $\leq \beta$ is safe if no MC with $v \in R$ exceeds β
- A region R for threshold $\leq \beta$ is unsafe if no MC with $v \in R$ is at most β

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Aim:

partition the parameter space into safe and unsafe regions

- Region = half-space defined by linear inequalities over the parameters
- A region R for threshold $\leq \beta$ is safe if no MC with $v \in R$ exceeds β
- A region R for threshold $\leq \beta$ is unsafe if no MC with $v \in R$ is at most β

We present two approaches:

- 1. An exact procedure.
- 2. An approximate technique.

How? Using SMT techniques How? Using parameter lifting

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Computing Rational Functions

[Daws, 2004]

Parameter Synthesis



Computing Rational Functions

[Daws, 2004]



Parameter Synthesis

 $\Pr\left(s_0 \models \diamondsuit(1 \text{ or } 3)\right) \leqslant \frac{1}{3} \quad \text{iff} \quad p \cdot q \cdot \frac{1-p}{1-p \cdot q} + p^2 \cdot \frac{1-q}{1-p \cdot q} \leqslant \frac{1}{3}$

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Computing Rational Functions

[Daws, 2004]



 $\Pr(s_0 \models \diamondsuit(1 \text{ or } 3)) \leqslant \frac{1}{3} \quad \text{iff} \quad p \cdot q \cdot \frac{1-p}{1-p \cdot q} + p^2 \cdot \frac{1-q}{1-p \cdot q} \leqslant \frac{1}{3}$

This may yield large high-degree rational functions.

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Resulting Rational Functions

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Zooming In

000000000000000000000000000vv16*x^10+11746794992985682921409606933 0000000000000000000000000000*v^11*x^14+2496820390874591873323652 00000*v'9*x'5+16230691765667091880343208999855273673085737156253814 I00000000*v^10*x^5+212270382782282965928362847764266275121524939197 0000000000000000000000v^11*x^10+94707045725196342033039966959729-138726376835263049363738601237408229827880859375000000000000000000

Zooming In

000000000000000000000000000vy^16*x^10+11746794992985682921409606933 00000*v*9*x*5+16230691765667091880343208999855273673085737156253814 /00000000*y^10*x^5+212270382782282965928362847764266275121524939197 94286762660590325279211306571960449218750000000000000000000000000 0000000000000000000000*y^11*x^10+94707045725196342033039966959729-138726376835263049363738601237408229827880859375000000000000000000

41 states, 138 transitions, 2 parameters:

numerator = 48 terms, denominator = product of 48 (linear) polynomials

Zooming In

100000000000000000000000000*v^8*x^16+7090864564569141868352890014 00000*v'9*x'5+16230691765667091880343208999855273673085737156253814 I000000001*v^10*x^5+212270382782282965928362847764266275121524939197 000000000000000000000*v^11*x^10+94707045725196342033039966959729-138726376835263049363738601237408229827880859375000000000000000000

41 states, 138 transitions, 2 parameters:

numerator = 48 terms, denominator = product of 48 (linear) polynomials

 \Rightarrow Use bisimulation, SCC-decomposition and efficient gcd-computation

Hierarchical SCC Decomposition

[Jansen *et al.*, 2014]



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[Jansen et al., 2014]



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Hierarchical SCC Decomposition



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[Jansen et al., 2014]



[Jansen et al., 2014]



For which (combinations of) values for p and q is the probability of reaching5smaller than $c \in [0, 1]$? \Rightarrow Evaluate rational function. Goal: partition parameter space in regions R that are either safe or unsafe Idea: generate region candidates R and ask SMT solver⁵ for counterexample

Exploiting SMT

Goal: partition parameter space in regions R that are either safe or unsafe Idea: generate region candidates R and ask SMT solver⁵ for counterexample



⁵Over non-linear real arithmetic using Z3 or SMT-RAT. < = > < = > < =



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For which $1/10 \le p \le 9/10$ and $2/5 \le q \le 3/5$ does $Pr(\diamondsuit 2) \ge 3/20$ hold?



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For which $1/10 \le p \le 9/10$ and $2/5 \le q \le 3/5$ does $Pr(\diamondsuit 2) \ge 3/20$ hold?



For which $1/10 \le p \le 9/10$ and $2/5 \le q \le 3/5$ does $Pr(\diamondsuit 2) \ge 3/20$ hold?

Experimental Results

competitors

- PARAM [Hahn et al., 2010]
- PRISM [Parker et al., 2011]

models

- Bounded retransmission protocol
- NAND multiplexing
- Zeroconf, Crowds protocol
- 10⁴ to 7.5 · 10⁶ states

experiments:

- best set-up for each tool
- log-scale x- and y-axis





Experimental Results

competitors

- PARAM [Hahn et al., 2010]
- PRISM [Parker et al., 2011]
- prototype [Baier et al., 2014]

models

- Bounded retransmission protocol
- NAND multiplexing
- Zeroconf, Crowds protocol
- 10⁴ to 7.5 · 10⁶ states

experiments:

- best set-up for each tool
- log-scale x- and y-axis

[Dehnert *et al.*, 2015]





runner-up in the CAV 2015 artefact evaluation

Parameter Synthesis using SMT

Pros:

- Exact results: rational function is an exact symbolic object
- Drastic improvements over existing tools

PARAM and PRISM

User-friendly representation

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Parameter Synthesis using SMT

Pros:

- Exact results: rational function is an exact symbolic object
- Drastic improvements over existing tools
 PARAM and PRISM
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Cons:

- Rational function requires many gcd-computations > 4 para
- SMT performance unpredictable

> 4 parameters? heuristics hard

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heuristics hard

Can we do better by sacrificing exactness? Yes.

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[Quatmann et al,, 2016]

Let transition probabilities be linear in each variable.

That is, transition functions f are multi-affine multivariate polynomials of form:

$$f = \sum a_i \cdot \left(\prod_{x \in V} x\right)$$
 with $a_i \in \mathbb{Q}$

Examples: $3x \cdot y + 4y \cdot z$, 1 - x, $x \cdot y \cdot z$ etc.

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[Quatmann et al,, 2016]

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[Quatmann et al,, 2016]



Parameter Synthesis

Two-phase approach: first remove dependencies, then substitute extremal values

[Quatmann et al,, 2016]



Parameter Synthesis

Two-phase approach: first remove dependencies, then substitute extremal values Also applicable to parametric MDPs.

Phase 1: Relaxation



Parameter dependencies are removed; $Pr(\diamondsuit 2) = (1-z) \cdot \frac{1-q}{1-p \cdot q}$ \Rightarrow each state is equipped with its own parameter

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Phase 1: Relaxation

Correctness:

- Relaxed regions contain more valuations than original regions
- \Rightarrow Relaxation yields over-approximations
- \Rightarrow Relaxation preserves upper-bounds on reachability probs

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Phase 1: Relaxation

Correctness:

- Relaxed regions contain more valuations than original regions
- \Rightarrow Relaxation yields over-approximations
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Complexity of parameter synthesis :

- Relaxation increases the number of parameters
- > Extremal values of the state parameters attain maximal probabilities
- \Rightarrow Valuations for maximal probabilities are easier to find

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Phase 2: Substitution



$$p, z \in [0.1, 0.3] \qquad p, z \in \{0.1, 0.3\} \\ q \in [0.4, 0.6] \qquad q \in \{0.4, 0.6\}$$

Local parameters per state \Rightarrow extremal values at states suffice

Joost-Pieter Katoen

Performance+Reliability by Model Checkin

109/131

Phase 2: Substitution



Local parameters per state \Rightarrow extremal values at states suffice

Phase 2: Substitution



This results in a Markov decision process.

Its extremal reachability probabilities provide bounds for parametric MC.

Parameter Synthesis



Until \approx 95% of the parameter space is covered

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Until 95% of the parameter space is covered

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Coverage

		φ	n	# states	# trans	%р	t	safe	unsafe	neither	unkn
pMC	brp	$\mathbb E$	2	20744	27 651	48%	51	14.9%	79.2%	5.8%	0.2%
		\mathbb{E}	4	20744	27 651	48%	71	7.5%	51.0%	40.6%	0.8%
	crowds	\mathbb{P}	2	104 512	246 082	19%	44	54.4%	41.1%	4.2%	0.3%
	nand	\mathbb{P}	2	35 112	52 647	47%	21	21.4%	68.5%	6.9%	3.2%
pMDP	brp	\mathbb{P}	2	40721	55 143	50%	153	6.6%	90.4%	3.0%	0.0%
	cons	\mathbb{P}	4	22 656	75 232	41%	357	2.6%	87.0%	10.4%	0.0%
	sav	\mathbb{P}	4	379	1 1 27	50%	2	44.0%	15.4%	35.4%	5.3%
	zconf	\mathbb{P}	2	88 858	203 550	40%	186	16.6%	77.3%	5.6%	0.5%

Parameter space $R = [10^{-5}, 1-10^{-5}]^n$ until 95% coverage for *n* parameters for 625 equally-sized regions without region refinement

single core, 2.0 GHz, 30GB RAM, TO = one hour

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Parametric Markov Chain Benchmarks

						PLA			PRISM
benchmark	instance	φ	#pars	#states	#trans	#regions	direct	bisim	best
	(256,5)	\mathbb{P}	2	19720	26 627	37	6	14	TO
	(4096,5)	\mathbb{P}	2	315 400	425 987	13	233	то	то
	(256, 5)	\mathbb{E}	2	20744	27 651	195	8	15	то
hua	(4096,5)	\mathbb{E}	2	331 784	442 371	195	502	417	то
brp	(16,5)	\mathbb{E}	4	1 304	1731	1251220	2764	1 597	то
	(32,5)	\mathbb{E}	4	2 600	3 459	1031893	то	2722	то
	(256,5)	\mathbb{E}	4	20744	27 651	-	то	то	то
	(10,5)	\mathbb{P}	2	104 512	246 082	123	17	6	2038
crowds	(15,7)	\mathbb{P}	2	8 364 409	25 108 729	116	1 880	518	то
	(20,7)	\mathbb{P}	2	45 421 597	164 432 797	119	то	2 935	то
nand	(10,5)	\mathbb{P}	2	35 112	52 647	469	22	30	TO
nanu	(25,5)	\mathbb{P}	2	865 592	1 347 047	360	735	2061	то

coverage of 95%; refinement into four equally-sized regions SMT approach needs >one hour on all instances.

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Parametric MDP Benchmarks

						PLA		PRISM	
benchmark	instance	φ	#pars	#states	#trans	#regions	direct	bisim	best
bra	(256,5)	\mathbb{P}	2	40 721	55 143	37	35	3 359	то
brb	(4096,5)	\mathbb{P}	2	647 441	876 903	13	3 4 2 4	ТO	то
	(2,2)	\mathbb{P}	2	272	492	119	< 1	< 1	31
conconcurs	(2,32)	\mathbb{P}	2	4 1 1 2	7 692	108	113	141	то
consensus	(4,2)	\mathbb{P}	4	22 656	75 232	6 125	1 866	2022	то
	(4,4)	\mathbb{P}	4	43 136	144 352	-	то	ΤO	то
	(6,2,2)	\mathbb{P}	2	379	1 127	162	< 1	< 1	TO
6317	(100, 10, 10)	\mathbb{P}	2	1 307 395	6 474 535	37	1 612	то	то
SdV	(6,2,2)	\mathbb{P}	4	379	1 127	621 175	944	917	то
	(10,3,3)	\mathbb{P}	4	1 850	6 561		то	ΤO	то
zoroconf	(2)	\mathbb{P}	2	88 858	203 550	186	86	1 295	TO
Zerocom	(5)	\mathbb{P}	2	494 930	1133781	403	2 400	ТO	то

coverage of 95%

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Summary So Far

SMT-based approach:

- Exact
- Requires rational functions
- Fickle SMT performance
- $\approx 10^6$ states, 2 parameters
- Restricted to Markov chains
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- Approximative
- Off-the-shelf model checking
- No SMT, no rational functions
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- Applicable to MDPs and games
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17/131

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Inputs:

- 1. a (finite) parametric MDP \mathcal{M} over $V = \{x_1, \ldots, x_n\}$ with signomial parameter functions $c \cdot x_1^{a_1} \cdot \ldots \cdot x_n^{a_n}$ for $c \in \mathbb{R}$
- 2. multiple objectives $\varphi_1, \ldots, \varphi_m$ (reachability, expected reward)
- 3. objective function f over V: $\sum_{k=1}^{N} c_k \cdot x_1^{a_{1k}} \cdot \ldots \cdot x_n^{a_{nk}}$ for $c_k \in \mathbb{R}$

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Output:

A (randomised) policy σ and valuation u such that:

$$\underbrace{\mathcal{M}^{\sigma}[u] \vDash \varphi_1 \land \ldots \land \varphi_m}_{\text{``feasibility''}} \quad \text{and} \quad \underbrace{\text{the objective } f \text{ is minimised}}_{\text{``optimality''}}$$

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multi-objective MDP: use LP [Etessami *et al.*, 2008] multi-objective parametric MDP: use special type NLP [Cubuktepe *et al.*, 2017]

Objectives: minimise f, reach T with probability $\leq p$, expected cost to reach $G \leq c$ Subject to: $p_{s_l} \leq p$ $c_{s_i} \leq c$ expected reward objective

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NLP for Two Objectives

 $\begin{array}{ll} \text{Objectives: minimise } f, \text{ reach } \mathcal{T} \text{ with probability } \leqslant p, \text{ expected cost to reach } G \leqslant c \\ \text{Subject to: } p_{s_l} \leqslant p & \text{reachability objective} \\ c_{s_l} \leqslant c & \text{expected reward objective} \\ \forall s: & \sum_{\substack{\alpha \in Act(s) \\ \forall s, \alpha : \end{array}} \sigma^{s, \alpha} = 1 & \text{randomised scheduler} \\ \forall s, \alpha : & 0 \leqslant \sigma^{s, \alpha} \leqslant 1 \end{array}$

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NLP for Two Objectives

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NLP for Two Objectives

Objectives: minimise f, reach T with probability $\leq p$, expected cost to reach $G \leq c$ reachability objective Subject to: $P_{s_l} \leq P$ $C_{S_i} \leq C$ expected reward objective $\forall s: \sum \sigma^{s,\alpha} = 1$ randomised scheduler $\forall s, \alpha : \quad 0 \leq \sigma^{s, \alpha} \leq 1$ $\forall s, \alpha : \sum_{t \in S} \mathcal{P}(s, \alpha, t) = 1$ probabilistic choice $\forall s, t, \alpha : \quad 0 \leq \mathcal{P}(s, \alpha, t) \leq 1$ $\forall s \in T : p_s = 1$ reach prob of T $\forall s \notin T: \quad p_s = \sum_{\alpha \in Act(s)} \sigma^{s,\alpha} \cdot \sum_{t \in S} \mathcal{P}(s,\alpha,t) \cdot p_t$ transition probabilities

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Theorem: This NLP is sound and complete. But solving NLPs is exponential.

Can We Do Better?



Yes.

- 1. Get a feasible solution in polynomial time⁶. How? Geometric programming.
- 2. Get local optimum. How? Sequential convex programming.

Solutions are approximations that can be arbitrarily close.

⁶Approximation of arbitrarily precise results by interior point methods with barriers Joost-Pieter Katoen Performance+Reliability by Model Checking 120/131

Geometric Programming

Objective: minimise
$$f :: \sum_{k=1}^{N} c_k \cdot x_1^{a_{1k}} \cdot \ldots \cdot x_n^{a_{nk}}$$
 for $c_k \in \mathbb{R}_{\ge 0}$
Subject to:
 $\forall i \in [1..m] : g_i \leq 1$ posynomial g_i
 $\forall j \in [1..\ell] : h_j = 1$ monomial h_j

Division transformation: $f \leq h$ if and only if $\frac{f}{h} \leq 1$

Relaxation: f = h implies $f \leq h$ if and only if $\frac{f}{h} \leq 1$

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Convexification

Lifting



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Objectives: reach T with probability $\leq p$, expected cost to reach $G \leq c$

Subject to:

$$\frac{P_{s_l}}{p} \leqslant 1$$
 reachability
 $\frac{C_{s_l}}{c} \leqslant 1$ expected reward

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Objectives: reach T with probability $\leq p$, expected cost to reach $G \leq c$

Subject to: $\frac{p_{51}}{1} \leq 1$

$$\forall s: \sum_{\substack{\alpha \in Act(s)\\ \sigma^{s,\alpha} \leqslant 1}} \sigma^{s,\alpha} \leqslant 1$$

reachability

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expected reward

randomised scheduler

Objectives: reach T with probability $\leq p$, expected cost to reach $G \leq c$

Subject to: $\frac{P_{s_l}}{p} \leq 1$

$$\begin{array}{c} & \overset{p}{\underset{c \in I}{s}} \\ & \overset{c_{s_{I}}}{\underset{c \in Act(s)}{s}} \leqslant 1 \\ \forall s : & \overset{\sigma}{\underset{\alpha \in Act(s)}{s}} \\ \forall s, \alpha : & \sigma^{s, \alpha} \leqslant 1 \\ \forall s, \alpha : & \sum_{\substack{t \in S}{\overline{\mathcal{P}}}(s, \alpha, t) \leqslant 1 \end{array} \end{array}$$
randomised scheduler

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reachability

Objectives: reach T with probability $\leq p$, expected cost to reach $G \leq c$

Subject to: $\frac{p_{s_l}}{p} \leq 1$ reachability <u>c_{s1}</u> ≤ 1 expected reward $\forall s: \sum \sigma^{s,\alpha} \leq 1$ randomised scheduler $\alpha \in Act(s)$ $\forall s, \alpha : \sigma^{s, \alpha} \leq 1$ $\forall s, \alpha : \sum_{t \in S} \overline{\mathcal{P}}(s, \alpha, t) \leq 1$ probabilistic choice $\forall s, t, \alpha : \overline{\mathcal{P}}(s, \alpha, t) \leq 1$ $\forall s \in T : p_s = 1$ reach prob of T $\forall s \notin T: \quad \underline{\sum_{\alpha} \sigma^{s, \alpha} \cdot \sum_{t \in S} \overline{\mathcal{P}}(s, \alpha, t) \cdot \underline{p_t}} \leq 1$ transition probabilities
GP for Two Objectives

Objectives: reach T with probability $\leq p$, expected cost to reach $G \leq c$

Subject to: $\frac{p_{s_l}}{p} \leq 1$ reachability <u>c_{s1}</u> ≤ 1 expected reward $\forall s: \sum \sigma^{s,\alpha} \leq 1$ randomised scheduler $\alpha \in Act(s)$ $\forall s. \alpha: \sigma^{s,\alpha} \leq 1$ $\forall s, \alpha : \sum_{t \in S} \overline{\mathcal{P}}(s, \alpha, t) \leq 1$ probabilistic choice $\forall s, t, \alpha : \quad \overline{\mathcal{P}}(s, \alpha, t) \leq 1$ $\forall s \in T : p_s = 1$ reach prob of T $\forall s \notin T: \quad \frac{\sum_{\alpha} \sigma^{s, \alpha} \cdot \sum_{t \in S} \mathcal{P}(s, \alpha, t) \cdot \mathbf{p}_t}{\sum_{\alpha} \sigma^{s, \alpha} \cdot \sum_{t \in S} \mathcal{P}(s, \alpha, t) \cdot \mathbf{p}_t} \leq 1$ transition probabilities $\forall s \notin G: \quad \frac{\sum_{\alpha} \sigma^{s, \alpha} \cdot \left(c(s, \alpha) + \sum_{t \in S} \overline{\mathcal{P}}(s, \alpha, t) \cdot c_t \right)}{\sum_{\alpha} \sigma^{s, \alpha} \cdot \left(c(s, \alpha) + \sum_{t \in S} \overline{\mathcal{P}}(s, \alpha, t) \cdot c_t \right)} \leq 1$ expected costs

Correctness

Use the objective function F now⁷

Minimise
$$\sum_{p \in V} \frac{1}{p} + \sum_{p \in L} \frac{1}{p} + \sum_{s,\alpha} \frac{1}{\sigma^{s,\alpha}}$$

yields that all variables p, \overline{p} and $\sigma^{s,\alpha}$ are maximised.

Theorem: The GP with objective function F yields a feasible solution. Solving this GP can be done in polynomial time.

⁷Note: the original objective function f is dropped.

Experimental Results

Benchmark	#states	#par	specs	MOSEK (s)		Z3
BRP (pMC)	5382	2	EC, ₽, *	23.17	(6.48)	-
	112646	2	$EC, \mathbb{P}, *$	3541.59	(463.74)	-
	112646	4	$EC, \mathbb{P}, *$	4173.33	(568.79)	
	5382	2	EC,\mathbb{P}	3.61		904.11
	112646	2	EC, \mathbb{P}	479.08		TO
NAND (pMC)	4122	2	$\mathrm{EC}, \mathbb{P}, *$	14.67	(2.51)	-
	35122	2	$EC, \mathbb{P}, *$	1182.41	(95.19)	-
	4122	2	EC,\mathbb{P}	1.25	di se cente	1.14
	35122	2	EC,\mathbb{P}	106.40		11.49
BRP (pMDP)	5466	2	EC, ₽, *	31.04	(8.11)	-
	112846	2	EC, ℙ, ∗	4319.16	(512.20)	-
	5466	2	EC,\mathbb{P}	4.93	******	1174.20
	112846	2	EC,\mathbb{P}	711.50		TO
CONS (pMDP)	4112	2	EC, ₽, *	102.93	(1.14)	-
	65552	2	$EC, \mathbb{P}, *$	TO	6 Y	-
	4112	2	EC,\mathbb{P}	6.13		TO
	65552	2	EC,\mathbb{P}	1361.96		TO

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Parameter Synthesis

Experimental Results



Joost-Pieter Katoen

Performance+Reliability by Model Checking

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Epilogue

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Geometric programming approach:

- Numerical approximation
- Multiple objectives
- $\approx 10^5$ states, 10 parameters
- Applicable to MDPs
- Possibility of richer objectives

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Significant progress in the last couple of years.

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Significant progress in the last couple of years.

More info: QEST'14, CAV'15, ATVA'16, TACAS'17 and

http://moves.rwth-aachen.de/research/tools/prophesy/

Overview

Epilogue

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Conclusion

Probabilistic Model Checking ...

- is a mature automated technique
- focuses on quantitative measures
- has a broad range of applications
- such as performance analysis
- and reliability engineering

more information: http://www.stormchecker.org

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Current Research

- tight game-based abstractions
- parameter synthesis

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Big Thanks! To all my co-authors and co-workers

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